Stochastic Hybrid Processes as Solutions of Stochastic Differential Equations

Estimation of Rare Event Probability in Stochastic Hybrid Systems

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Outline

- Introduction
- Classical SDE models
- Stochastic Hybrid Systems
 - Model by Ghosh and Bagchi
 - Models by Krystul and Blom
- Estimation of rare event probability in stochastic hybrid systems

Stochastic Differential Equation

$$dX_t = a(X_t) dt + \sigma(X_t) dW_t, \quad t > 0$$
$$X(0) = X_0$$

 $\{W_t, t \ge 0\}$ is the standard Brownian motion. $a(\cdot)$: drift coefficient and $b(\cdot)$: diffusion coefficient

Solution : existence and uniqueness; under "Lipschitz" and "Growth" condition

(L)
$$|a(x) - a(y)|^2 + |b(x) - b(y)|^2 \le K|x - y|^2$$

(G) $|a(x)|^2 + |b(x)|^2 \le K|x|^2$

SDE: Strong solution: Properties

1. Continuous path

2. Semimartingale : X(t) = X₀ + A(t) + M(t) A(t) is of bounded variation, M(t) is a Martingale. - Important for mathematical analysis (well developed theory)

3. <u>Markov</u>: $\mathcal{D}(X(t)|\mathcal{F}_s) = \mathcal{D}(X(t)|X(s)), \quad t > s.$ Transition prob. : $p(s, x, t, A) = P(X(t) \in A | X(s) = x)$ Time homogeneous : p(s, x, t, A) = p(0, x, t - s, A) $\Rightarrow \mathcal{D}(X(t+s)|\mathcal{F}_s) = \mathcal{D}_{X(s)}(X(t)).$

Strong Markov :

 $\mathcal{D}\left(X(t+\tau)|\mathcal{F}_{\tau}\right) = \mathcal{D}_{X(\tau)}\left(X(t)\right), \ \tau \text{ stopping time} \\ - \text{Helpful to analyse processes stopped at random times}$

Jump Diffusion Process

- Extension while keeping Semimartingale property?
- Discontinuous paths?
- Jump Diffusion Process :

 $dX_t = a(X_t) dt + b(X_t) dW_t + h(X_t) dN_t, \quad t > 0$ {N(t)}: (pure) jump process.

Alternative representation: $dX_t = a(X_t) dt + b(X_t) dW_t + \int_{\mathbb{R}} g(X_{t-}, u) p(dt, du)$

 $p(\cdot, \cdot)$ Poisson random measure associated with the jump process.

Solution: existence, uniqueness under Growth and Lipschitz condition on a, b and g.

Properties : Semimartingale, Markov, but Discontinuous

Stochastic Hybrid Systems

Further extension : Not only the process but also the governing model/equation jumps.

- regime switch or Hybrid model.

Model I (Ghosh & Bagchi):

$$(X_t, \theta_t) \in \mathbb{R}^d \times \mathbb{M}, \ \mathbb{M} = \{1, \dots, N\}, \ t \ge 0$$

$$dX_{t} = a(X_{t}, \theta_{t}) dt + b(X_{t}, \theta_{t}) dW_{t} + \int_{\mathbb{R}} g(X_{t-}, \theta_{t-}, u) p(dt, du)$$
$$P(\theta_{t+\delta t} = j | \theta_{t} = i, X_{s}, \theta_{s}, s \leq t) = \lambda_{ij}(X_{t}) \delta t + o(\delta t), \quad i \neq j$$
$$X(0) = X_{0} \quad \theta(0) = \theta_{0}$$

 $p(\cdot, \cdot)$ — Poisson random measure with intensity $dt \times l(du)$ $\lambda_{ij}(\cdot) \ge 0, i \ne j, i, j = 1, 2, ... N$ and $\sum_{j=1}^{N} \lambda_{ij}(\cdot) = 0$

Assumptions :

- $a(\cdot,i), \ b(\cdot,i)$ are bounded and Lipschitz continuous
- $\lambda_{ij}(\cdot)$ are bounded and measurable
- Support of $g(x, \theta, u)$ w.r.t. 'u' (i.e., the proj. on \mathbb{R}) is bdd.

Existence of Solution: (Put in Ito-Skorohod's framework) Identify i with e_i , ith unit vector in \mathbb{R}^N and embed \mathbb{M} into \mathbb{R}^N . For $x \in \mathbb{R}^d$, define $\Delta_{ij}(x)$ to be the consecutive (with lexicographic ordering on $(i, j) \in \mathbb{M} \times \mathbb{M}$) left closed right open intervals on \mathbb{R}_+ with length $\lambda_{ij}(x)$.

Define $c : \mathbb{R}^d \times \mathbb{M} \times \mathbb{R} \to \mathbb{R}^N$ as $c(x, i, u) = e_j - e_i \quad \text{if} \quad u \in \Delta_{ij}(x), \quad 0 \text{ otherwise.}$

Model can be expressed as

 $dX_t = a(X_t, \theta_t) dt + b(X_t, \theta_t) dW_t + \int_{\mathbb{R}} g(X_{t-}, \theta_{t-}, u) p(dt, du)$

$$d\theta_t = \int_{\mathbb{R}} c(X_{t-}, \theta_{t-}, u) \, p(dt, du), \quad t > 0$$

$$X(0) = X_0 \quad \theta(0) = \theta_0$$

and the existence of unique strong solution can be shown.

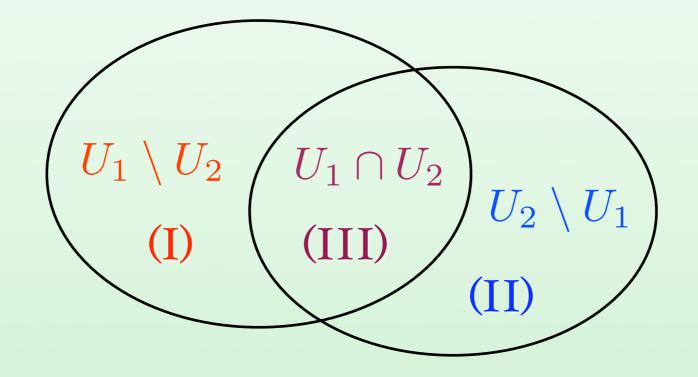
Remarks :

- Conditions on the coefficients can be relaxed
- $\{X(t)\}$ is <u>not</u> Markov
- Augmented process $(X(t), \theta(t))$ is Markov

Let U_1 and U_2 denote respectively the supports of $g(x, \theta, u)$ and $c(x, \theta, u)$ w.r.t. 'u' (i.e. the projections on \mathbb{R}).

Different situations can arise:

- I. Process jumps but no switching: $u \in U_1 \setminus U_2$
- II. No jump in process but switching occurs: $u \in U_2 \setminus U_1$
- III. Simultaneous switching and jump (Hybrid Jump): $u \in U_1 \cap U_2$



Hybrid Systems - Model 2 (Krystul and Blom)

$$dX_{t} = a(X_{t}, \theta_{t}) dt + b(X_{t}, \theta_{t}) dW_{t} + \int_{\mathbb{R}^{n}} g_{1}(X_{t-}, \theta_{t-}, u) q_{1}(dt, du) + \int_{\mathbb{R}^{n}} g_{2}(X_{t-}, \theta_{t-}, u) p_{2}(dt, du)$$

$$d\theta_t = \int_{\mathbb{R}^n} c(X_{t-}, \theta_{t-}, u) \, p_2(dt, du), \quad t > 0$$

- W is an m-dimensional standard Wiener process.
- $q_1(dt, du)$ is a martingale random measure associated to a Poisson random measure p_1 with intensity $dt \times m_1(du)$.
- $p_2(dt, du)$ is a Poisson random measure with intensity $dt \times m_2(du) = dt \times du_1 \times \mu(\underline{u}), \mu$ is a probability measure on $\mathbb{R}^{n-1}, u_1 \in \mathbb{R}.$

Assumptions :

- Lipschitz condition. For all $i=1,2,\ldots,N$

$$a(x,e_i)^2| + |b(x,e_i)|^2 + \int_{\mathbb{R}^n} |g_1(x,e_i,u)|^2 m_1(du) \le K(1+|x|^2).$$

• Growth condition:

$$a(x, e_i) - a(y, e_i)|^2 + |b(x, e_i) - b(x, e_i)|^2 + \int_{\mathbb{R}^n} |g_1(x, e_i, u) - g_1(x, e_i, u)|^2 m_1(du) \le K_r |x - y|^2.$$

for $|x| \leq r, |y| \leq r$.

- $\lambda_{ij}(\cdot)$ are bounded and measurable
- Support of $g_2(x, \theta, u)$ w.r.t. ' u_1 ' (i.e., the proj. on \mathbb{R}) is bdd.

Hybrid Systems - Model 3 (Krystul, Blom and Bagchi)

Idea:

 X_t starts in some set E^i and (X_t, θ_t) evolves according to some Ito-Skorohod SDE.

If X_t reaches a boundary of that set, the process immediately jumps to a new set E^j (possibly including jump in θ)

Process continues again according to the Ito-Skorohod SDE in a new set until it reaches the boundary of the current set.

$$E = \{x \mid x \in E^{i}, \text{ for some } i = 1, \dots, L\} = \bigcup_{i=1}^{L} E^{i}$$
$$\partial E = \{x \mid x \in \partial E^{i}, \text{ for some } i = 1, \dots, L\} = \bigcup_{i=1}^{L} \partial E^{i}$$

 $(X_t, \theta_t) \in \mathbb{R}^d \times \mathbb{M}, \quad \mathbb{M} = \{e_1, e_2, \dots, e_N\}.$ Suppose $\tau_1^E < \tau_2^E < \dots < \tau_m^E < \dots$ are the successive hitting times of ∂E .

Between the boundary jumps $\{X_t, \theta_t\}$ is a switching jump-diffusion process. For $\tau_m^E < t < \tau_{m+1}^E$,

$$\begin{aligned} X_t &= X_{\tau_m^E} + \int_{\tau_m^E}^t a(X_s, \theta_s) ds + \int_{\tau_m^E}^t b(X_s, \theta_s) dW_s \\ &+ \int_{\tau_m^E}^t \int_{\mathbb{R}^n} g_2(X_{s-}, \theta_{s-}, u) p_2(ds, du), \end{aligned}$$
$$\theta_t &= \theta_{\tau_m^E} + \int_{\tau_m^E}^t \int_{\mathbb{R}^n} c(X_{s-}, \theta_{s-}, u) p_2(ds, du). \end{aligned}$$

At jump points $(t = \tau_m)$,

$$X_{\tau_m^E} = f^x (X_{\tau_m^E}, \theta_{\tau_m^E}, Z),$$

$$\theta_{\tau_m^E} = f^\theta (X_{\tau_m^E}, \theta_{\tau_m^E}, Z),$$

where Z is some V-valued random variable,

$$f^{x}: \partial E \times \mathbb{M} \times V \longrightarrow E,$$
$$f^{\theta}: \partial E \times \mathbb{M} \times V \longrightarrow \mathbb{M}.$$

Assumptions:

- a and b satisfy Lipschitz and Growth conditions
- $\lambda_{ij}(\cdot)$ are bounded and measurable
- Support of $g_2(x, \theta, u)$ w.r.t. ' u_1 ' (i.e., the proj. on \mathbb{R}) is bdd.
- Function g_2 has the following property: $(x + g_2(x, \theta, u)) \in E$ for each $x \in E^i$, $\theta \in \mathbb{M}$, $u \in \mathbb{R}^n$, $i = 1, \dots, L$
- $d(\partial E, f^x(\partial E, \mathbb{M}, V)) > 0$,

i.e. when $\{X_t\}$ has reached the boundary ∂E it always jumps inside of open set E.

Estimation of rare event probability in stochastic hybrid systems

 $\{X_t, \theta_t\} \in \mathbb{R}^n \times \mathbb{M}, \quad \mathbb{M} = \{e_1, \dots, e_S\}$ switching diffusion process:

 $dX_t = a(\theta_t, X_t)dt + b(\theta_t, X_t)dW_t,$ $P(\theta_{t+\delta} = \theta | \theta_t = \eta, X_t = x) = \lambda_{\eta\theta}(x)\delta + o(\delta), \eta \neq \theta.$

 (X_t, θ_t) starts at t = 0 in $D_0 \subset \mathbb{R}^n \times \mathbb{M}$ with known initial probability distribution $P_{X_0, \theta_0}(\cdot)$.

 $\tau_A = \inf\{t \ge 0 : X_t \in A\}$: hitting time of $A \subset \mathbb{R}^n, A \cap D_0 = \emptyset$.

To calculate $P(\tau_A < T)$: probability that X_t will reach the set A in the time interval (0, T).

Using crude Monte Carlo to estimate $P(\tau_A < T)$ may require a huge number of simulations. For events with probability of the order 10^{-10} , often 10^{11} or more Monte Carlo simulated trajectories are needed.

Factorization Approach

<u>Idea</u> : Identify intermediate states that are (sequentially) visited much more often than the rare target set.

 D_0 initial set;

 $D_1 \supset \cdots \supset D_{m-1} \supset D_m = A \subset \mathbb{R}^n; \quad D_0 \cap D_1 = \emptyset.$ Define for $k = 1, \ldots, m$, stopping times $\tau_k \triangleq \inf\{t \ge 0 : X_t \in D_k\}$

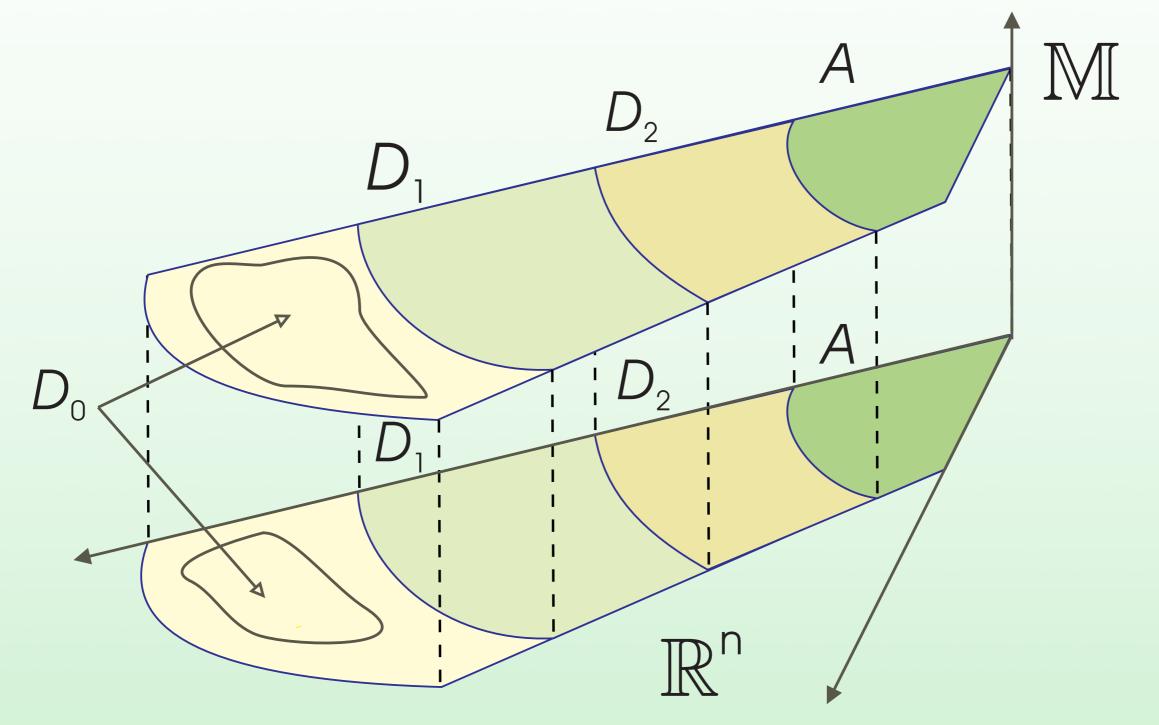
$$P(\tau_A < T) = P(\tau_m < T)$$

= $P(\tau_m < T, \tau_{m-1} < T, \dots, \tau_0 < T)$
= $\prod_{k=1}^{m} P(\tau_k < T | \tau_{k-1} < T).$

Individual probabilities on r.h.s. are not very small.

Nested sequence of level sets

 $D_0 \text{ initial set; } D_1 \supset \cdots \supset D_{m-1} \supset D_m = A; \quad D_0 \cap D_1 = \emptyset.$ $P(\tau_A < T) = \prod_{k=1}^m P(\tau_k < T | \tau_{k-1} < T).$



Factorization Approach (continued) $P(\tau_A < T) = \prod_{k=1}^{m} P(\tau_k < T | \tau_{k-1} < T) \triangleq \prod_{k=1}^{m} \gamma_k$ To calculate γ_k 's we proceed as follows:

Define
$$\xi_k = (X_{\tau_k \wedge T}, \theta_{\tau_k \wedge T}),$$

 $\pi_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_k < T), \quad B \subset \mathbb{R}^n \times \mathbb{M},$
 $p_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_{k-1} < T)$

 $\{(X_t, \theta_t)\}$ strong Markov $\Rightarrow \{\xi_k\}$ is a Markov (chain)

Factorization Approach (continued) $P(\tau_A < T) = \prod_{k=1}^{m} P(\tau_k < T | \tau_{k-1} < T) \triangleq \prod_{k=1}^{m} \gamma_k$ To calculate γ_k 's we proceed as follows: Define $\xi_k = (X_{\tau_k \wedge T}, \theta_{\tau_k \wedge T}),$ $\pi_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_k < T), \quad B \subset \mathbb{R}^n \times \mathbb{M},$ $p_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_{k-1} < T)$ $\{(X_t, \theta_t)\}$ strong Markov $\Rightarrow \{\xi_k\}$ is a Markov (chain) $p_k(B) = \int_{\mathbb{D}^n \times \mathbb{M}} P_{\xi_k | \xi_{k-1}}(B|\xi) \pi_{k-1}(d\xi),$ $\pi_k(B) = \frac{P(\xi_k \in B, \tau_0 < T, \dots, \tau_k < T)}{P(\tau_0 < T, \dots, \tau_k < T)} = \frac{P(\xi_k \in B \cap D_k, \tau_0 < T, \dots, \tau_{k-1} < T)}{P(\tau_0 < T, \dots, \tau_k < T)}$ $= \frac{P(\xi_k \in B \cap \bar{D}_k | \tau_0 < T, \dots, \tau_{k-1} < T)}{P(\tau_k < T | \tau_0 < T, \dots, \tau_{k-1} < T)} = \frac{\int_B^{T} \mathbf{I}_{\{\xi \in \bar{D}_k\}} p_k(d\xi)}{\int_{\mathbb{R}^n \times \mathbb{M}}^{T} \mathbf{I}_{\{\xi \in \bar{D}_k\}} p_k(d\xi)}$ $\gamma_k = \int_{\mathbb{R}^n \times \mathbb{M}} \mathbf{I}_{\{\xi \in \bar{D}_k\}} p_k(d\xi)$ $D_k \triangleq D_k \times \mathbb{M}$

Factorization Approach (continued) $P(\tau_A < T) = \prod_{k=1}^{m} P(\tau_k < T | \tau_{k-1} < T) \triangleq \prod_{k=1}^{m} \gamma_k$ To calculate γ_k 's we proceed as follows:

Define
$$\xi_k = (X_{\tau_k \wedge T}, \theta_{\tau_k \wedge T}),$$

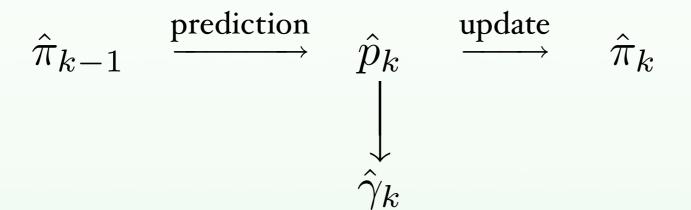
 $\pi_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_k < T), \quad B \subset \mathbb{R}^n \times \mathbb{M},$
 $p_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_{k-1} < T)$

 $\{(X_t, \theta_t)\}$ strong Markov $\Rightarrow \{\xi_k\}$ is a Markov (chain) Evolution of the flow $\{\pi_k, p_k, \gamma_k; k = 0, 1, \dots, m\}$ is described by the following diagram:

with initial condition $\pi_0(d\xi) = P_{\xi_0}(d\xi) = P(\xi_0 \in d\xi)$.

Interacting Particle System (IPS) approach

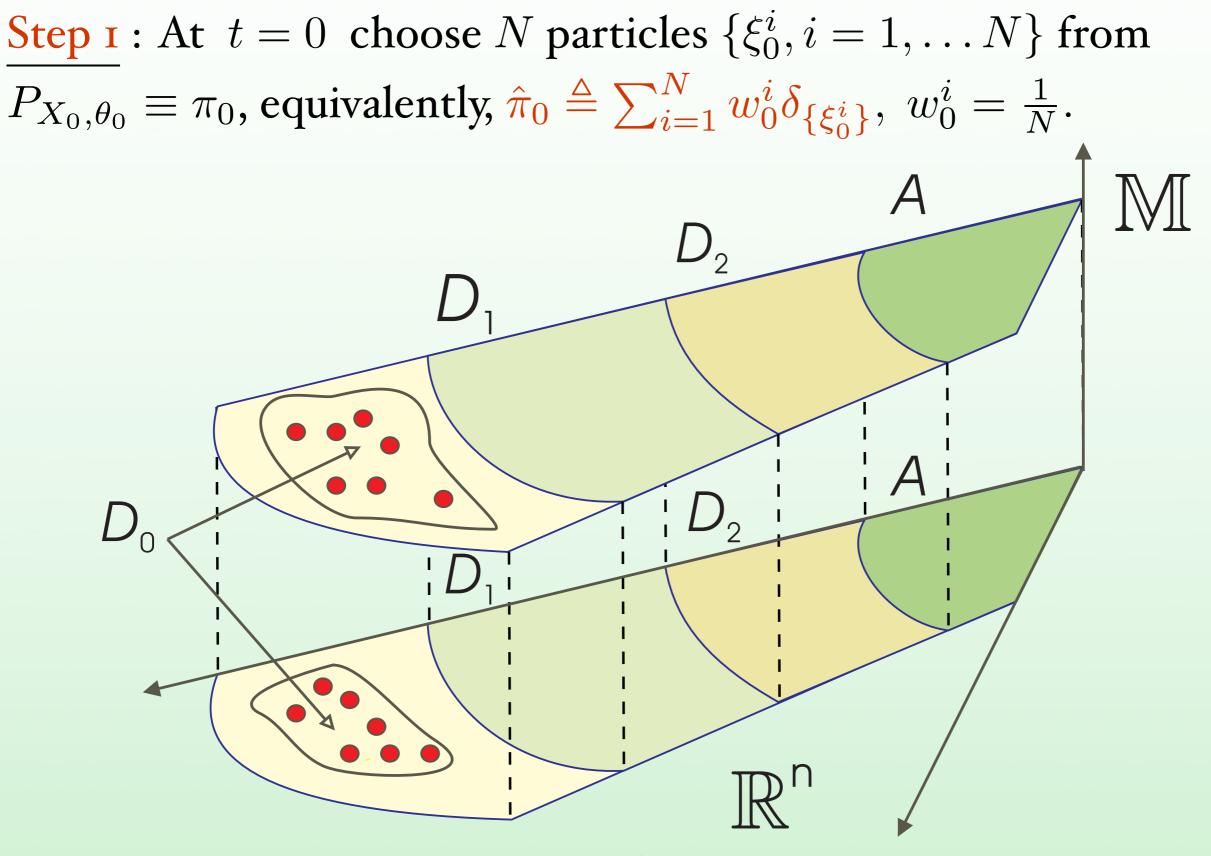
An approximating sequence $\{\hat{\pi}_k, \hat{p}_k, \hat{\gamma}_k; k = 0, 1, \dots, m\}$:



Approximations in the form of weighted empirical distributions associated with the particle system $\{\xi_k^i, \omega_k^i\}_{i=1}^N$:

$$\pi_{0} \approx \hat{\pi}_{0} = \sum_{i=1}^{N} \omega_{0}^{i} \delta_{\{\xi_{0}^{i}\}}, \quad \pi_{k} \approx \hat{\pi}_{k} = \sum_{i=1}^{N} \frac{\mathbf{I}_{\{\xi_{k}^{i} \in \bar{D}_{k}\}}}{\sum_{j=1}^{N} \omega_{k}^{j} \mathbf{I}_{\{\xi_{k}^{j} \in \bar{D}_{k}\}}} \delta_{\{\xi_{k}^{i}\}},$$
$$p_{k} \approx \hat{p}_{k} = \sum_{i=1}^{N} \omega_{k}^{i} \delta_{\{\xi_{k}^{i}\}}, \quad \gamma_{k} \approx \hat{\gamma}_{k} = \sum_{i=1}^{N} \omega_{k}^{i} \mathbf{I}_{\{\xi_{k}^{i} \in \bar{D}_{k}\}}.$$

IPS algorithm



Step 2. Prediction step: $\pi_{k-1} \rightarrow p_k$

Given $\hat{\pi}_{k-1} \triangleq \sum_{i=1}^{N} w_{k-1}^i \delta_{\{\xi_{k-1}^i\}}$, i.e., a weighted particle cloud $\{(\xi_{k-1}^i, w_{k-1}^i), i = 1, \dots, N\}$ with $\xi_{k-1}^i \in \overline{D}_{k-1}$,

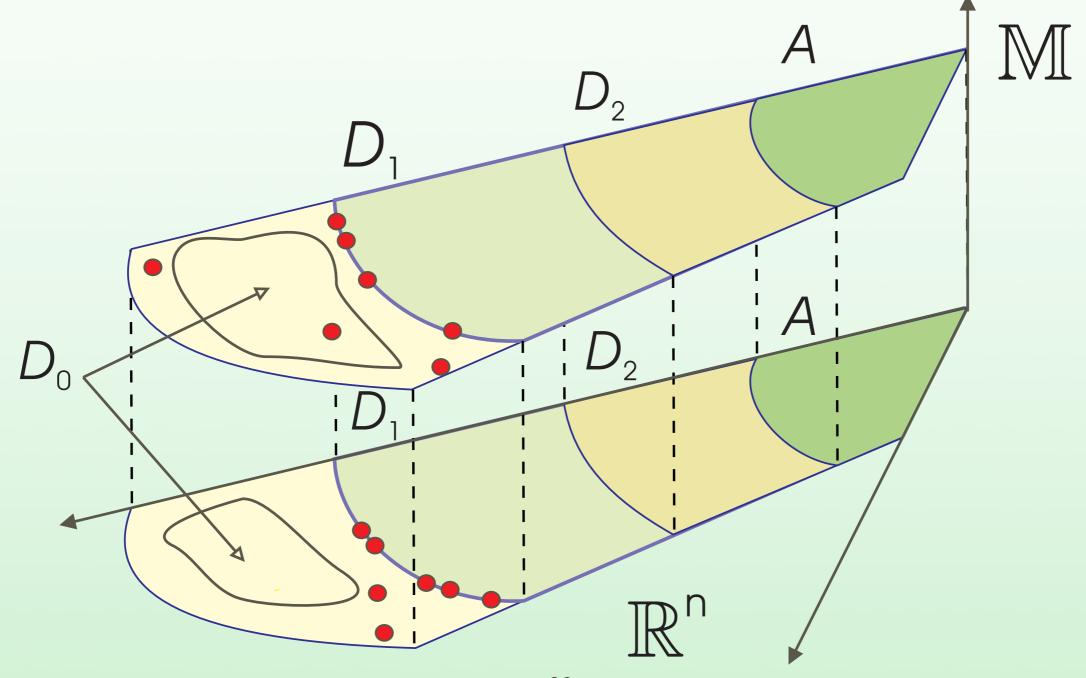
Let particles evolve following the hybrid system equations until \overline{D}_k or the final time T is hit.

 $\hat{\xi}_k^i$: Value of the *i*-th particle at the end of the step $\hat{p}_k = \sum_{i=1}^N w_{k-1}^i \delta_{\{\hat{\xi}_k^i\}}$ is the approximation of p_k .

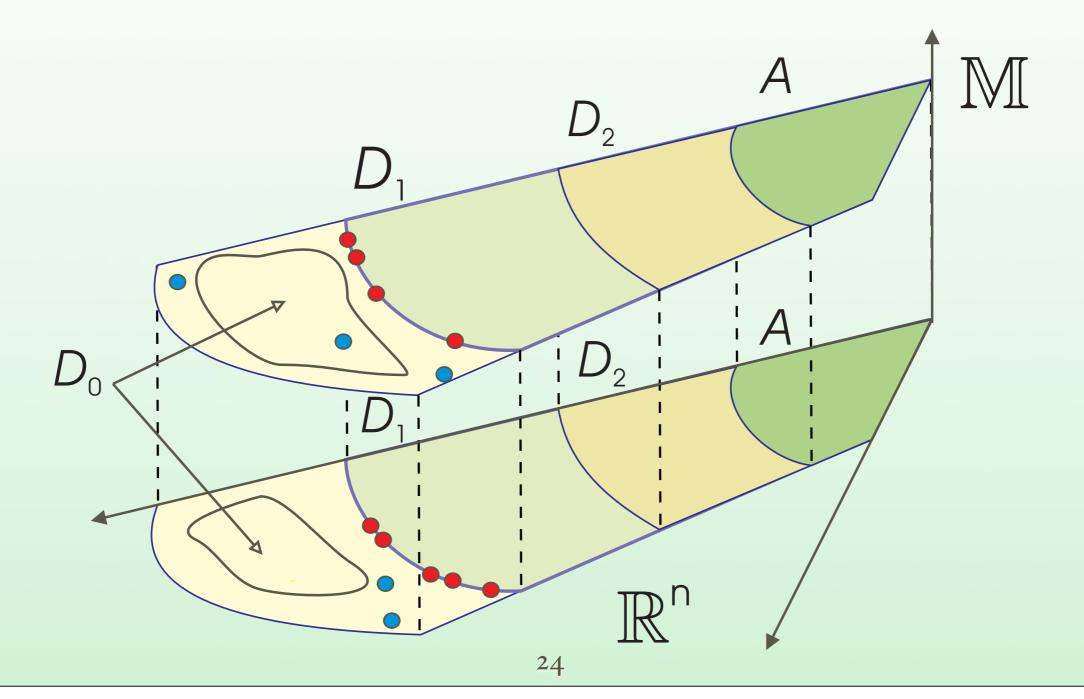
Recall that:

$$p_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_{k-1} < T), B \subset \mathbb{R}^n \times \mathbb{M}.$$

 $\hat{\xi}_k^i$: Value of the *i*-th particle at the end of the prediction step $\hat{p}_k = \sum_{i=1}^N w_{k-1}^i \delta_{\{\hat{\xi}_k^i\}}$ is the approximation of p_k .



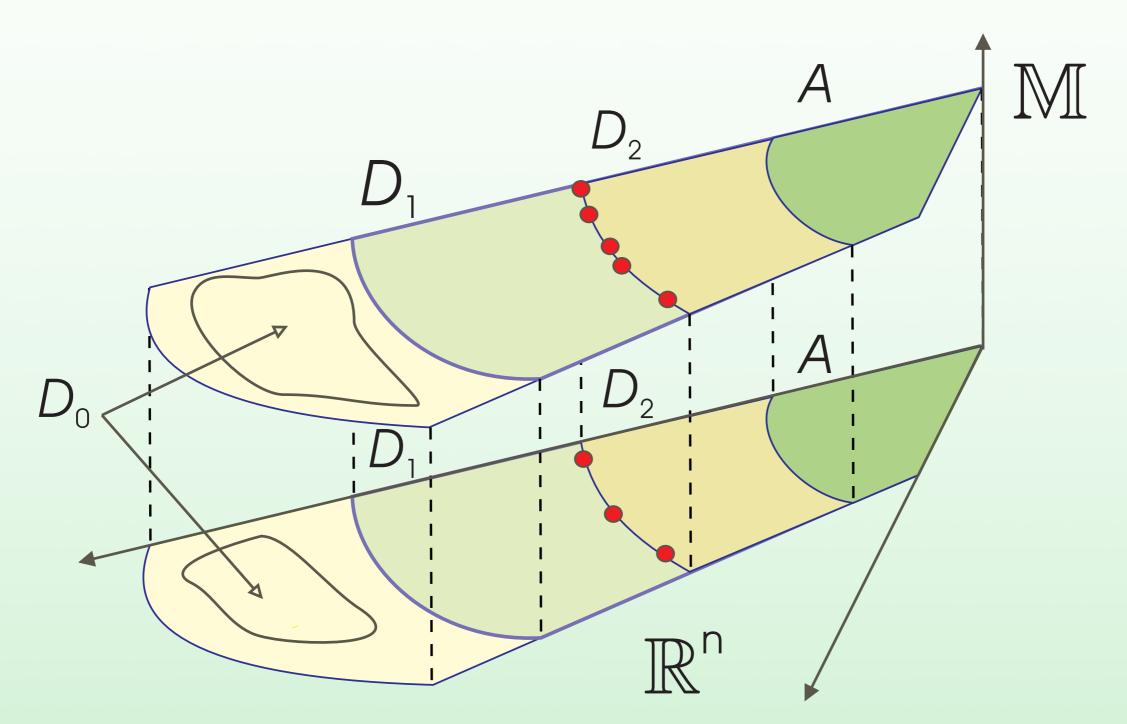
Set $\hat{w}_k^i = 0$, if $X_{\tau_k \wedge T}^i \notin D_k$, otherwise $\hat{w}_k^i = w_{k-1}^i$. Approximate $\gamma_k = P(\tau_k < T | \tau_{k-1} < T)$ by $\hat{\gamma}_k = \sum_{i=1}^N \hat{w}_k^i$. Stop algorithm if $\hat{\gamma}_k = 0$, and $P(\tau_A < T) \approx \prod_{k=1}^m \hat{\gamma}_k = 0$.



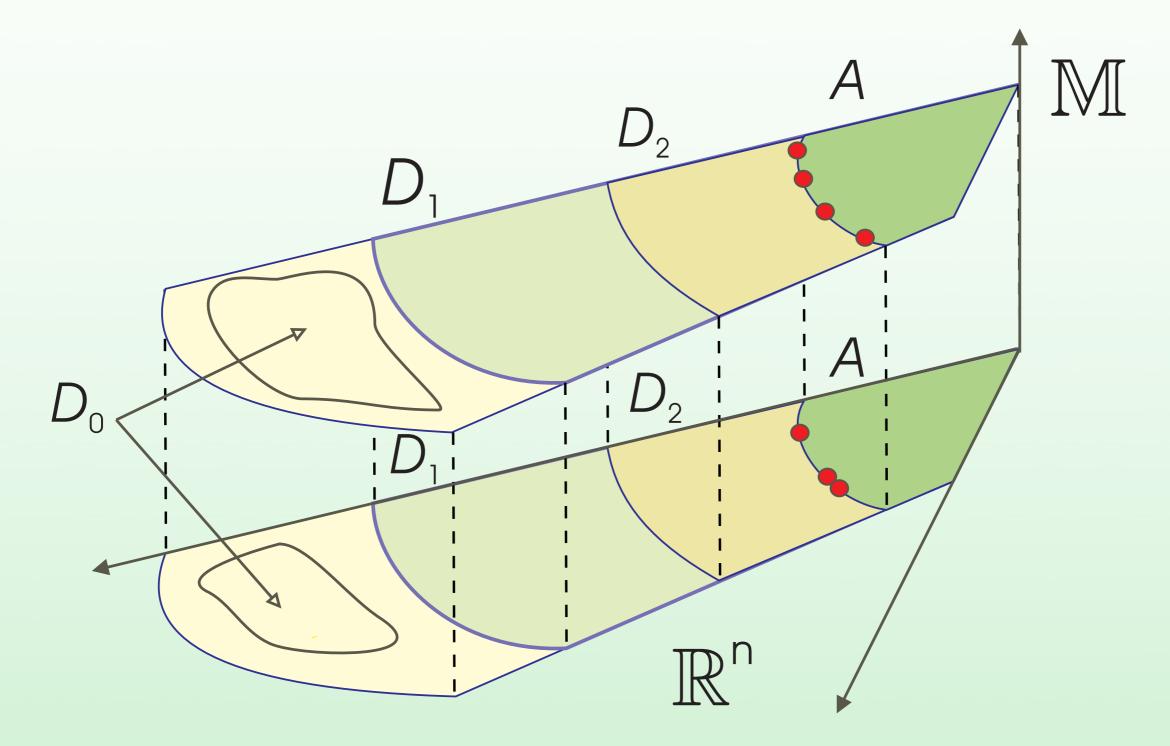
Step 3. Updating step: $p_k \to \pi_k$ Approximate π_k by $\hat{\pi}_k \triangleq \sum_{i=1}^N \tilde{w}_k^i \delta_{\{\hat{\xi}_k^i\}}$, with $\tilde{w}_k^i = \hat{w}_k^i / \sum_{i=1}^N \hat{w}_k^i$. Recall that: $\pi_k(B) = P(\xi_k \in B | \tau_0 < T, \dots, \tau_k < T), B \subset \mathbb{R}^n \times \mathbb{M}$. To avoid carrying the particles with no weight, resample Nparticles from $\hat{\pi}_k$. The new set of particles is $\{\xi_k^i, \omega_k^i\}_{i=1}^N$ with $w_k^i = \frac{1}{N}$.

Go to Step 2.

k := k + 1







Theorem (Cerou, Del Moral, LeGland and Lezaud, 2002)

IPS estimator is unbiased, i.e.

$$\mathbb{E}\Big[\prod_{k=1}^{m} \gamma_k^N\Big] = P(\tau_A < T) = P_{hit}(0,T)$$

and

$$\left(\mathbb{E}\left(\prod_{k=1}^{m}\gamma_{k}^{N}-\prod_{k=1}^{m}\gamma_{k}\right)^{p}\right)^{\frac{1}{p}} \leq \frac{a_{p}b_{m}}{\sqrt{N}},$$

for some finite constant a_p which depends only on the parameter p, and for some finite constant b_m which depends only on the parameter m.

Diffusion Example: Geometric Brownian Motion

$$dX_t = (\mu + \frac{\sigma^2}{2})X_t dt + \sigma X_t dW_t, \quad X_0 = x$$

Probability of hitting level d before time T : $P(\tau_d \leq T), \quad \tau_d \triangleq \inf\{t > 0 : X_t \geq d\}.$

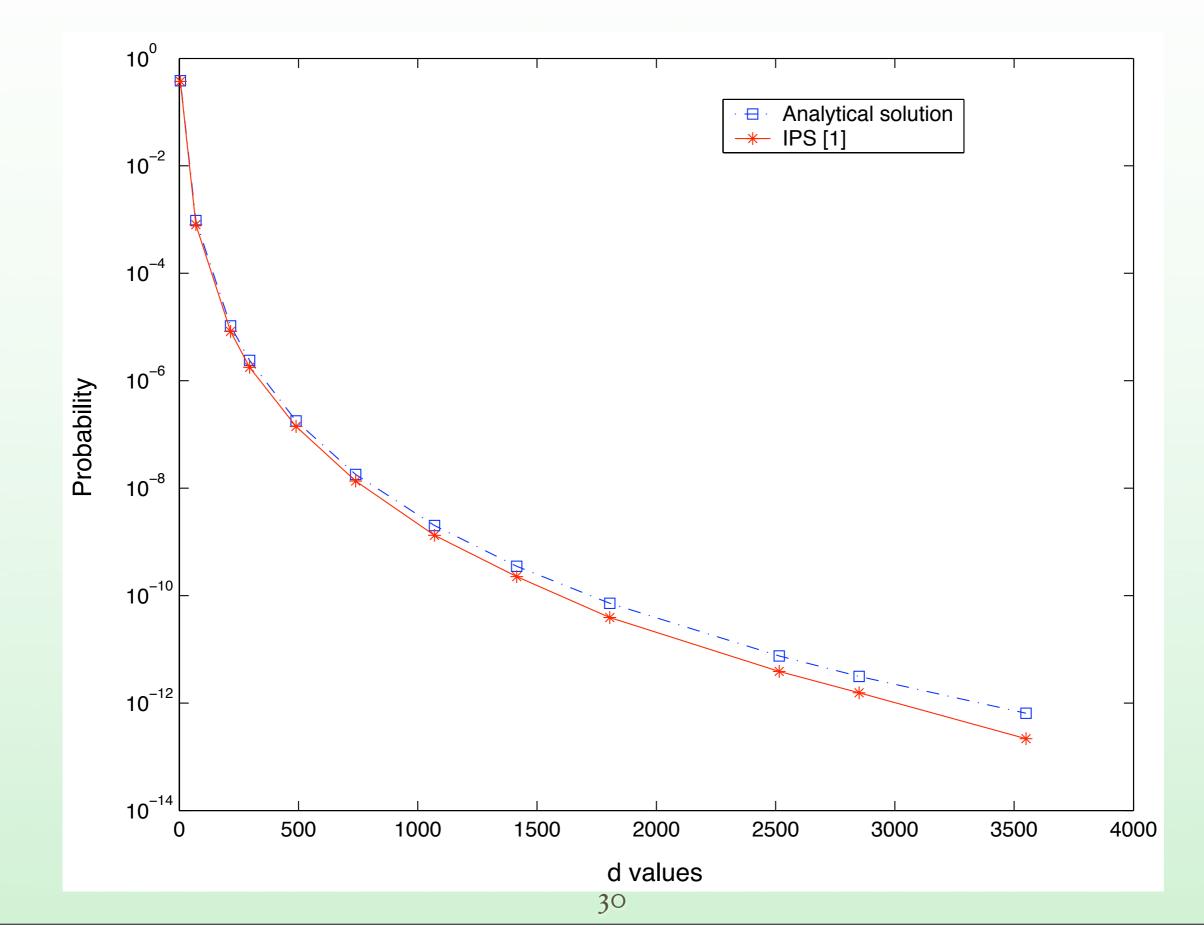
Analytical solution:

$$\int_{0}^{t} \frac{\ln(d/x)}{\sqrt{2\pi\sigma^{3}s^{3}}} \exp\{\frac{-(\ln(d/x) - \mu s)^{2}}{2\sigma^{2}s}\} ds$$

\underline{IPS} :

- $\mu = 1, \sigma = 1, x_0 = 1, \text{ varying } d \le 3550$
- intermediate levels d_j 's are chosen experimentally; 40%-55% of particles starting at D_{j-1} reaches D_j .
- 1000 simulations of 1000 particles each $P_{hit}(0,T) \approx \frac{1}{1000} \sum_{i=1}^{1000} (\prod_{k=1}^{m} \hat{\gamma}_k)^{(i)}$

Probability to hit level d before time T = 1: diffusion



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